

# Math & Stat for MBA

## Problem Set #1

Student ID  
Sungchen Park

Ans 1-1 Where  $y_t = \beta_1 + \beta_2 x_t + \epsilon_t$ ,  $t = 1, 2, \dots, T$   
The least squares estimator  $\hat{\beta} = (X'X)^{-1}X'y$

$$i) X'X = \begin{bmatrix} 1 & \dots & 1 \\ x_1 & \dots & x_T \end{bmatrix} \begin{bmatrix} 1 & x_1 \\ \vdots & \vdots \\ 1 & x_T \end{bmatrix} = \begin{bmatrix} T & \sum_t x_t \\ \sum_t x_t & \sum_t x_t^2 \end{bmatrix}$$

$$(X'X)^{-1} = \begin{bmatrix} \sum_t x_t^2 & -\sum_t x_t \\ -\sum_t x_t & T \end{bmatrix} / \Delta, \quad \text{where } \Delta = T \sum_t x_t^2 - (\sum_t x_t)^2 \\ = T \sum_t (x_t - \bar{x})^2$$

$$ii) X'y = \begin{bmatrix} 1 & \dots & 1 \\ x_1 & \dots & x_t \end{bmatrix} \begin{bmatrix} y_1 \\ \vdots \\ y_t \end{bmatrix} = \begin{bmatrix} \sum_t y_t \\ \sum_t x_t y_t \end{bmatrix}$$

Then, we can carry out  $\hat{\beta}$

$$\hat{\beta} = (X'X)^{-1}X'y = \begin{bmatrix} \sum_t x_t^2 & -\sum_t x_t \\ -\sum_t x_t & T \end{bmatrix} \begin{bmatrix} \sum_t y_t \\ \sum_t x_t y_t \end{bmatrix} \times \frac{1}{T \sum_t (x_t - \bar{x})^2} \\ = \begin{bmatrix} \sum_t x_t^2 \cdot \sum_t y_t - \sum_t x_t \cdot \sum_t x_t y_t \\ -\sum_t x_t \cdot \sum_t y_t + T \sum_t x_t y_t \end{bmatrix} \times \frac{1}{T \sum_t (x_t - \bar{x})^2} = \begin{bmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{bmatrix}$$

$$\therefore \hat{\beta}_1 = \bar{y} - \frac{\sum_t (x_t - \bar{x}) y_t}{\sum (x_t - \bar{x})^2} \quad \hat{\beta}_2 = \frac{\sum_t (x_t - \bar{x}) y_t}{\sum (x_t - \bar{x})^2}$$

Ans 1-2

As we know,  $\text{Cov}(\hat{\beta}) = \sigma^2 (X'X)^{-1}$

$$= \sigma^2 \times \frac{1}{T \sum_t (\lambda_t - \bar{\lambda})^2} \begin{bmatrix} \sum_t \lambda_t^2 & -\sum_t \lambda_t \\ -\sum_t \lambda_t & T \end{bmatrix}$$

$$= \frac{\sigma^2}{\sum_t (\lambda_t - \bar{\lambda})^2} \begin{bmatrix} \sum_t \lambda_t^2 / T & -\bar{\lambda} \\ -\bar{\lambda} & 1 \end{bmatrix} = \begin{bmatrix} \text{Var}(\hat{\beta}_1) & \text{Cov}(\hat{\beta}_1, \hat{\beta}_2) \\ \text{Cov}(\hat{\beta}_1, \hat{\beta}_2) & \text{Var}(\hat{\beta}_2) \end{bmatrix}$$

Ans 2-1

Assump that the electricity industry follows CRS (Constant Returns to Scale).  
Then, Output increases by the same proportional change as all inputs changes.  
In this case, Cost represents all inputs.

Cost =  $A \cdot \text{Output}^\alpha$ ,  $\alpha$  must be 1 to satisfy assumption.

if  $\alpha > 1$  Decreasing RS.

$\alpha < 1$  Increasing RS.

i)  $H_0: \beta_2 = 1$  (CRS)

$H_1: \beta_2 < 1$  (IRS)

ii)  $t = \frac{\hat{\beta}_2 - \beta_2}{\text{S.e}(\hat{\beta}_2)} = \frac{0.720 - 1}{0.0175} = -16 < -t_{(140, 5\%)} = -1.655$



thus, we reject  $H_0$

Ans 2-2

Where  $Y = A^\alpha B^\beta C^\gamma$ , if function  $Y$  is Homogeneous of Degree one in  $A, B, C$   
then,  $\alpha + \beta + \gamma = 1$

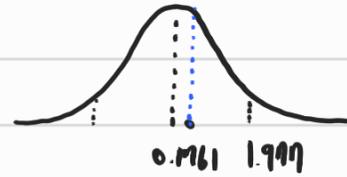
i)  $H_0: \beta_3 + \beta_4 + \beta_5 = 1$

$H_1: \beta_3 + \beta_4 + \beta_5 \neq 1$

$$ii) t = \frac{\beta_3 + \beta_4 + \beta_5 - 1}{S.E(\hat{\beta}_3 + \hat{\beta}_4 + \hat{\beta}_5)}, \quad \text{Var}(\hat{\beta}_3 + \hat{\beta}_4 + \hat{\beta}_5) = \text{Var}(\hat{\beta}_3) + \text{Var}(\hat{\beta}_4) + \text{Var}(\hat{\beta}_5) + 2(\text{Cov}(\hat{\beta}_3, \hat{\beta}_4) + \text{Cov}(\hat{\beta}_3, \hat{\beta}_5) + \text{Cov}(\hat{\beta}_4, \hat{\beta}_5))$$

$$= 0.1761 < t_{(140, 2.5\%)} = 1.9977$$

then, Do not reject  $H_0$



Ans 3-1

$$\hat{y}_0 = X_0' \hat{\beta} \quad \text{denote } 0 \text{ as the number of observations}$$

first, we should find  $\hat{\beta}$

$$i) \hat{\beta} = (X'X)^{-1} X'y = \begin{bmatrix} \frac{2}{3} & -\frac{1}{3} \\ -\frac{1}{3} & \frac{2}{3} \end{bmatrix} \begin{bmatrix} 2 \\ 1 \end{bmatrix} = \begin{bmatrix} 1 \\ 0 \end{bmatrix} = \begin{bmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{bmatrix}$$

$$ii) \hat{\sigma}^2 = \frac{1}{n-2} \sum \hat{e}_i^2 = \frac{1}{9} \sum (y_i - \hat{y})^2 = \frac{1}{9} \sum (y_i - \hat{\beta}_1 x_{1i} - \hat{\beta}_2 x_{2i})^2$$

$$= \frac{1}{9} \sum (y_i - x_{1i})^2 = \frac{1}{9} \left\{ \frac{11}{3} - 4 + 2 \right\} = \frac{1}{27}$$

$$iii) \hat{y}_{12} = X_{12}' \hat{\beta} = [5 \ -2] \begin{bmatrix} 1 \\ 0 \end{bmatrix} = 5$$

$$S.E(\hat{y}_{12}) = \hat{\sigma} \sqrt{X_{12}' (X'X)^{-1} X_{12} + 1} = \frac{1}{\sqrt{27}} \times \sqrt{27} = 1$$

then, we can define 80% prediction intervals as

$$\therefore \hat{y}_{12} \pm t_{(9, 10\%)} \cdot S.E(\hat{y}_{12}) = 5 \pm 1.383 \cdot 1 = [3.617, 6.383]$$

iv) As same as for  $\hat{y}_{12}$

$$\hat{y}_{13} = 3, \quad S.E(\hat{y}_{13}) = \frac{1}{\sqrt{27}} \times \sqrt{\frac{161}{3}} = 1.41$$

$$\therefore \hat{y}_{13} \pm t_{(9, 10\%)} \cdot S.E(\hat{y}_{13}) = 3 \pm 1.383 \cdot 1.41 = [3.05, 6.95]$$

Ans 3-2

for the expected value of  $y_{12}, y_{13}$ , only change the variances.

$\mu_{12}$  is the mean of the line at  $x_{12}$

i)  $\mu_{12}$  has the same value of  $\hat{y}_{12} = 5$

$$\text{ii) } S.E(\mu_{12}) = \hat{\sigma} \sqrt{X'_{12}(X'X)^{-1}X_{12}} = \frac{1}{\sqrt{27}} \times \sqrt{26} = 0.98$$

80% prediction intervals is then,

$\mu_{12}$

$$\mu_{12} \pm t_{(9, 10\%)} \cdot S.E(\mu_{12}) = 5 \pm 1.383 \cdot 0.98 = [3.64, 6.36]$$

iii) As same as for  $\mu_{13}$

$$\mu_{13} \pm t_{(9, 10\%)} \cdot S.E(\mu_{13}) = 3 \pm 1.383 \cdot 1.182 = [3.31, 6.63]$$

Ans 3-3

At Q 3-1, the answer represents a particular point  $\hat{y}_{12}, \hat{y}_{13}$

On the other hands, the answer of Q 3-2 predicts the mean of the line at  $x_{12}, x_{13}$

Unlike the mean of the line, a particular point inevitably has extra noise term  $\epsilon$ .

So, it brings more variance when we find prediction intervals

That's the reason why we got two different answer at the same circumstance.

